

# Models for Predicting the Development of Insulin Resistance

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**Abstract:** Insulin resistance is the leading cause for developing type 2 diabetes. Early determination of insulin resistance and herewith of impending type 2 diabetes could help to establish sooner preventive measures or even therapies. However, an optimal predictive model for developing insulin resistance has not been established yet. Based on the data of an Austrian cohort study (SAPHIR study) various predictive models were calculated and compared to each other. For developing predictive models logistic regression models were used. For finding an optimal cut-off value an ROC approach was used. Based on various biochemical parameters an overall percentage of around 82% correct classifications could be achieved.

## 1 INTRODUCTION

Insulin resistance (IR) is a condition in which, muscle, fat and liver cells do not respond properly to insulin and these cells therefore cannot easily absorb glucose from the bloodstream. So, the body needs higher levels of insulin to help glucose enter the cells and as consequence the so called beta cells in the pancreas subsequently increase their production of insulin to try to keep up with this increased demand for insulin by producing more insulin. As long as the beta cells are able to produce enough insulin to overcome the insulin resistance of the cells, blood glucose levels stay in a healthy range. But over time the beta cells fail to keep up with the body's increased need for insulin and without enough insulin, excess glucose builds up in the bloodstream. This can lead to diabetes and other serious health disorders (Rutter et al., 2008).

Diabetes is often asymptomatic and around 50% of patients suffering from type 2 diabetes are undiagnosed and the delay from disease onset to diagnosis can exceed 10 years. At the time patients are diagnosed around 25% have established retinopathy and around 50% have signs of diabetic tissue damage (Griffin et al., 2000). Furthermore type 2 diabetes is affecting around 8% of the current global adult population and the prevalence is growing worldwide (Keating, 2015).

Early determination of insulin resistance and herewith of impending type 2 diabetes could help to establish sooner preventive measures or even therapies.

The gold-standard for quantifying insulin resistance is the hyperinsulinemic euglycemic clamp but this impractical in epidemiological studies because the measurement takes about two hours and medical attention is needed. Therefore a commonly used surrogate endpoint is the homeostatic model assessment (HOMA) which is derived from the product of fasting glucose and insulin levels.

The HOMA index is a robust tool for the assessment of insulin resistance in epidemiological studies. HOMA index values larger 2.0 are used as an indication for insulin resistance (Griffin et al., 2000).

The aim of this work was to develop a predictive model for insulin resistance based on various clinical parameters and to evaluate this model based on an ROC approach.

## 2 DATA

As study population the data of an Austrian cohort study (SAPHIR study, Salzburg Atherosclerosis Prevention program in subjects at High Individual Risk) was used.

The prospective SAPHIR study was conducted from 1999 to 2002 and involved unrelated male and female subjects between 40 and 70 years, who lived in the greater Salzburg region and responded to invitations by their family or workplace physician or to announcement in the local press. The study has been approved by the ethical committee of Salzburg and written consent was granted by the participants.

At baseline and after approximately 3 years the participants were subjected to a screening program that included a personal and family history and a physical examination. These evaluations also included a panel of laboratory tests, measurement of visceral fat mass, body composition, insulin sensitivity, intima-media thickness of the carotid arteries and a detailed medical history. HOMA indices have been used to define the level of insulin sensitivity.

In total, 1,770 subjects were included in the SAPHIR study. All subjects of the SAPHIR study were invited to a second follow-up examination 5 years after their completion of the SAPHIR study.

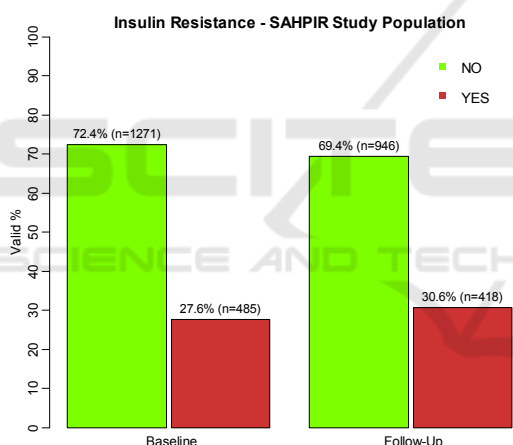


Figure 1: Insulin resistance, SAPHIR study population.

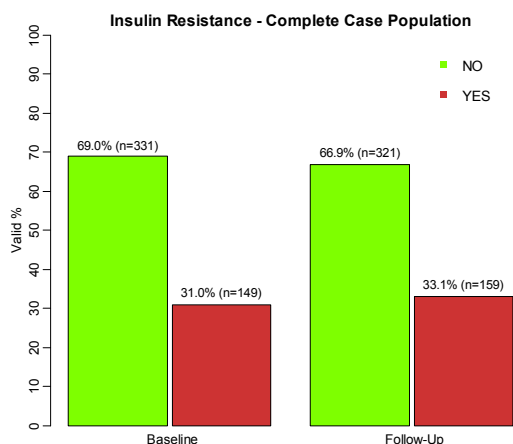


Figure 2: Insulin resistance, complete case population.

This resulted in a response of 1386 subjects, which were used for the model building process. Because of missing data only 480 (“complete case population”) of these 1386 subjects could be used.

The distribution of insulin resistance was similar between the two populations. The largest difference was 3.4 percent points. Subjects with insulin resistance at baseline were also included into the model because in “real life” the status of insulin resistance of the subject at baseline is unknown, so exclusion would lead to a potential bias

### 3 PREDICTIVE MODELS

#### 3.1 Approach

Because of the binary nature of the data a logistic regression approach was chosen to predict the target variable “indication of insulin resistance at follow-up”. The baseline covariates there chosen from a clinical point of view.

Table 1: Continuous baseline covariates.

Baseline covariates	No IR (mean ± SD)	IR (mean ± SD)	p-value
2hr-OGTT glucose level	32.38 ± 28.98	81.47 ± 64.68	<0.001**
Adiponectin [µg/ml]	9.19 ± 4.74	6.94 ± 3.67	<0.001**
Age [years]	51.06 ± 6.08	52.26 ± 5.72	0.001**
BMI [kg/m <sup>2</sup> ]	25.55 ± 3.33	30.09 ± 4.20	<0.001**
Fat mass [kg]	16.82 ± 7.8	25.13 ± 9.26	<0.001**
F-Insulin [µU/ml]	5.14 ± 1.82	13.53 ± 5.98	<0.001**
Glucose [mg/dl]	89.17 ± 8.79	104.91 ± 27.71	<0.001**
HbA1c [%]	5.53 ± 0.30	5.87 ± 0.86	<0.001**
HDL cholesterol [mg/dl]	62.73 ± 15.71	51.43 ± 12.37	<0.001**
HOMA-Index	1.14 ± 0.42	3.53 ± 1.97	<0.001**
kITT-Index	4.51 ± 1.20	3.24 ± 1.17	<0.001**
Lean mass [kg]	59.35 ± 11.65	64 ± 12.96	<0.001**
Triglyceride [mmol/l]	108.7 ± 74.21	168.27 ± 102.39	<0.001**
Waist-Hip-Ratio	0.88 ± 0.08	0.93 ± 0.07	<0.001**

Table 2: Categorical baseline covariates.

Baseline covariates	No IR (% , n)	IR (% , n)	p-value
Gender			
female	476 (72.8%)	178 (27.2%)	0.783
male	795 (72.1%)	307 (27.9%)	
Hypertension			
no	789 (80.8%)	188 (19.2%)	<0.001**
yes	457 (62.0%)	280 (38.0%)	
Incident diabetes			
no	1261 (75.0%)	421 (25.0%)	<0.001**
yes	10 (13.5%)	64 (86.5%)	

Hypertension was defined as a systolic blood pressure above 130 mmHg and a diastolic blood pressure above 85 mmHg. In the early stages of diabetes sometimes no insulin resistance has

developed yet, therefore baseline incident diabetes was also included as a covariate in the model. Incident diabetes was defined as a fasting plasma glucose level  $\geq 126$  mg/dl, a 2-hr OGTT glucose of  $\geq 200$  mg/dl or current use of hypoglycaemic drug therapy regardless the reason.

Continuous data was compared between the two groups using the two sample independent t-test in case of normality (verification with the Kolmogorov-Smirnov test with Lilliefors correction) and variance homogeneity (verification with the Levene test). If continuous data was normally distributed but variance was heterogeneous Welch's t-test was used and if data showed no normal distribution, the Mann-Whitney-U test was used. Unpaired categorical data was compared using Fisher's exact test. For calculations SPSS (IBM Corp. Released 2013. IBM SPSS Statistics for Windows, Version 22.0. Armonk, NY: IBM Corp.) and the statistical computing software R Version 3.2.3 (R Foundation for Statistical Computing, Vienna, Austria. URL <http://www.R-project.org>) were used. The uncorrected type I error was set at 5% (two-sided), this means no adjustment for the type I error was made.

### 3.2 First Model

The a-priori selected baseline covariates (except gender) discriminated very well regarding the indication of baseline insulin resistance, so for a first model all baseline covariates were used to predict indication of insulin resistance at follow-up.

The estimates based on this logistic regression model (model Ia) are presented in the table below.

Table 3: Variables of the model Ia.

Baseline Covariates	p-value	Odds-Ratio	95% CI for OR
2hr-OGTT glucose level [mg/dl]	0.001**	1.015	1.006-1.024
Adiponectin [ $\mu$ g/ml]	0.038*	0.911	0.834-0.995
Age [years]	0.547	0.983	0.931-1.039
BMI [ $\text{kg}/\text{m}^2$ ]	0.077	1.169	0.983-1.39
Fat mass [kg]	0.407	0.972	0.91-1.039
F-Insulin [ $\mu$ U/ml]	0.001**	1.706	1.24-2.348
Glucose [mg/dl]	0.011*	1.058	1.013-1.106
HbA1c [%]	0.977	0.986	0.397-2.452
HDL cholesterol [mg/dl]	0.010*	0.963	0.936-0.991
HOMA-Index	0.008**	0.212	0.067-0.672
kITT-Index	0.893	1.017	0.794-1.303
Lean mass [kg]	0.980	1.000	0.951-1.051
Triglyceride [mmol/l]	0.979	1.000	0.996-1.004
Waist-Hip-Ratio	0.394	9.645	0.052-1773.282
Gender (male vs. female)	0.037*	3.886	1.086-13.897
Hypertension (no vs. yes)	0.395	0.776	0.432-1.393
Incident diabetes (no vs. yes)	0.708	1.403	0.238-8.265
Constant	0.004**	0.000	1.006-1.024

This model yielded an overall percentage of 82.2% correct classifications (90.5% correct classifications for no indication of insulin resistance and 65.3% correct classifications for indication of insulin resistance). A Nagelkerke's  $R^2$  of 0.550 indicated also a good model fit.

Based on the model an increase in one unit of "2hr-OGTT glucose level [mg/dl]", "F-Insulin [ $\mu$ U/ml]" and "Glucose [mg/dl]" is associated with a significant higher chance for insulin resistance. An increase of one unit in "Adiponectin [ $\mu$ g/ml]" and "HDL cholesterol [mg/dl]" is associated with a significant lower chance for insulin resistance. There was also a significant increased chance for women suffering from insulin resistance at follow-up.

Interestingly low levels of HOMA-index baseline were associated with a higher chance of insulin resistance at follow-up (Odds Ratio = 0.212). But a univariate analysis yielded the expected association (Odds Ratio = 4.315,  $p < 0.001^{**}$ ). So this unexpected multivariate Odds Ratio may be caused through the multicollinearity of the model.

The confidence interval of the "Waist-Hip-Ratio" was extremely large so a bootstrapping approach based on 1000 bootstrap-samples was used to validate the numerical integrity of model Ia. The confidence interval for "Waist-Hip-Ratio" was even more extreme so this variable was excluded from the model. The results from this model (model Ib) remained almost the same like model Ia.

The classification results of the model Ib are presented in the table below.

Table 4: Classification result of model Ib.

Observed	Predicted		% correct
	no IR	IR	
	272	29	90.4
	IR	51	65.5
Overall %			82.2

(Cut-Off value 0.5)

### 3.3 Second Model

The first model yielded a good overall percentage of correct classification results but was relatively complex because various laboratory parameters were needed. So the next step was to reduce the complexity of the model using a variable selection approach (Bursac et al., 2008). A reduced model with similar results but fewer needed laboratory parameters would also be more cost efficient.

The first approach was using a backwards-likelihood-ratio-approach (probability for entry 0.05, probability for removal 0.10) which led to an overall percentage of correct classification of 81.7% but this

model was again relatively complex (8 covariates were selected). So the next step was using a forwards-likelihood-ratio-approach (probability for entry 0.05, probability for removal 0.10) which resulted in a similar overall percentage of correct classifications but with a reduced number of covariates.

The estimates based on this forward-logistic regression model (model II) are presented in the table below.

Table 5: Variables of the model II.

Baseline Covariates	p-value	Odds-Ratio	95% CI for OR
2hr-OGTT glucose level [mg/dl]	0.001**	1.013	1.005-1.021
Adiponectin [µg/ml]	0.012*	0.900	0.83-0.977
BMI [kg/m <sup>2</sup> ]	0.002**	1.123	1.044-1.207
F-Insulin [µU/ml]	<0.001**	1.188	1.095-1.289
HDL cholesterol [mg/dl]	0.001**	0.959	0.936-0.982
Gender (male vs. female)	0.005**	2.599	1.337-5.052
Constant	0.001**	0.019	1.005-1.021

The result was a relatively compact model with an overall percentage of 81.9% correct classifications (91.6% correct classifications for no indication of insulin resistance and 62.3% correct classifications for indication of insulin resistance). A Nagelkerke’s R<sup>2</sup> of 0.531 indicated again a good model fit.

Based on this model an increase in one unit of “2hr-OGTT glucose level [mg/dl]”, “BMI [kg/m<sup>2</sup>]” and one unit of “F-Insulin [µU/ml]” is associated with a significant higher chance for insulin resistance.

An increase in one unit of “Adiponectin [µg/ml]” and one unit of “HDL cholesterol [mg/dl]” is associated with a significant lower chance for insulin resistance. There was also a significant increased chance for women suffering from insulin resistance at follow-up.

The robustness of this model was again verified with a bootstrapping approach based on 1000 bootstrap-samples; no inconsistency was detected.

Compared to the overall percentage of correct classifications and correct classifications for no indication of insulin resistance the percentage of correct classifications for indication of insulin resistance was low. So based on a Receiver Operating Characteristic curve (ROC curve) an optimal cut-off value for the predicted probabilities was searched.

#### 4 ROC ANALYSIS

A standard technique for summarizing the performance of a predictive model over a range of

cut-off-values is the Receiver Operating Characteristic curve (ROC curve) (Sweets, 1988). Since the ROC curve summarizes the predictive power of a model over all possible cut-off-values, it is more informative than a simple classification table for a fixed cut-off value.

The area under the curve (AUC) can be used as a performance metric for ROC curves. A higher area under the curve indicates a better prediction power. The previous logistic regression models yielded an area under the curve of 0.896 (model Ib) respectively 0.890 (model II).

For the previous logistic regression models the standard cut-off-value of 0.5 was used for the predicted probabilities. To maximize the effectiveness of a predictive model the cut-off value can be varied to achieve a higher true positive rate (“sensitivity”) and a higher true negative rate (“specificity”). The Youden index method (Youden, 1950) was used to define the optimal cut-off value for the two models.

The idea of the Youden index is to maximize the difference between true positive rate and false positive rate (“1 – specificity”). This means in other words that the Youden index is specified as the point

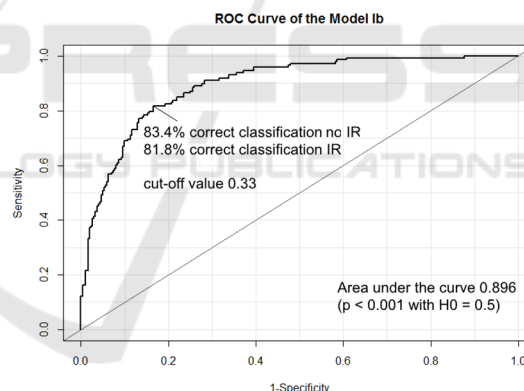


Figure 3: ROC curve for model Ib.

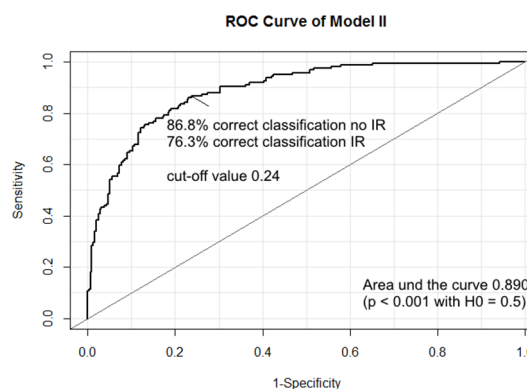


Figure 4: ROC curve for model II.

on the ROC curve which has the largest distance from the diagonal line. Based on the Youden index an optimal cut-off-value of 0.33 (model Ib) respectively 0.24 (model II) could be found.

In other papers an achieved area under the curve of 0.71 (Chiang et al., 2011) or 0.75 (Behboudi-Gandevani et al., 2016) was reported. In (Er et al., 2016) various models for predicting insulin resistance with an achieved area under the curve from 0.31 up to 0.80 were reported. So the area under the curve of the two presented predictive models indicates a slightly better performance.

## 5 CONCLUSIONS

Based on the data of an Austrian cohort study a predictive model with an overall percentage of 81.9% correct classifications for predicting insulin resistance based on gender and 5 biochemical parameters could be developed. Furthermore the classification cut-off-value was optimized by the use of ROC curves to achieve 86.6% correct classifications for no indication of insulin resistance and 76.3% correct classifications for indication of insulin resistance. Additionally based on the R package “caret” (Kuhn, 2008) a random forest approach was used for predicting insulin resistance. This approach yielded a similar overall percentage (79.2%) of correct classifications (86.6% correct classifications for no indication of insulin resistance and 64.2% correct classifications for indication of insulin resistance).

Limitations of this model are that the data was collected in a relatively small area and that because of various missing values in the baseline covariates and target variable only around 480 of 1386 patients could be included in the model building process. To account for this relatively large percentage of missing values a multiple imputation approach (5 imputations, fully conditional specification method based on an iterative Markov chain Monte Carlo algorithm) was used. The results of this approach yielded very similar results except that the influence of the covariate gender was not significant anymore (Odds Ratio = 0.997;  $p = 0.989$ ) like the univariate test of gender between baseline insulin resistance.

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